Herschel-ATLAS: The link between accretion luminosity and star formation in quasar host galaxies

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ABSTRACT

We use the science demonstration field data of the Herschel-ATLAS to study how star formation, traced by the far-infrared Herschel data, is related to both the accretion luminosity and redshift of quasars selected from the Sloan Digital Sky Survey and the 2SLAQ survey. By developing a maximum likelihood estimator to investigate the presence of correlations between the far-infrared and optical luminosities we find evidence that the star-formation in quasar hosts is correlated with both redshift and quasar accretion luminosity. Assuming a relationship of the form \( L_{IR} \propto L_{QSO}^\theta (1+z)^\zeta \), we find \( \theta = 0.22 \pm 0.08 \) and \( \zeta = 1.6 \pm 0.4 \), although there is substantial additional uncertainty in \( \zeta \) of order \( \pm 1 \), due to uncertainties in the host galaxy dust temperature. We find evidence for a large intrinsic dispersion in the redshift dependence, but no evidence for intrinsic dispersion in the correlation between \( L_{QSO} \) and \( L_{IR} \), suggesting that the latter may be due to a direct physical connection between star formation and black hole accretion. This is consistent with the idea that both the quasar activity and star formation are dependent on the same reservoir of cold gas, so that they are both affected by the influx of cold gas during mergers or heating of gas via feedback processes.

Key words: galaxies: active – galaxies: high-redshift – quasars: general
1 INTRODUCTION

AGN activity is now widely believed to be an important phase in the evolution of every massive galaxy in the Universe. This belief stems firstly from the discovery of a relatively tight correlation between the mass of a galaxy’s bulge and the mass of its central supermassive black hole (SMBH; e.g. Magorrian et al. 1998; Ferrarese & Merritt 2000; Gebhardt et al. 2000), which implies that build-up of the stellar mass and the SMBH mass are causally connected (although see also Jahnke & Maccìò 2010, who show that averaging of parameters by mergers can also produce such a correlation).

From a more theoretical perspective, AGN have become a key ingredient for semi-analytic models of galaxy formation (e.g. Benson et al. 2003; Cole et al. 2000; Guiderdoni et al. 1998; Granato et al. 2001). The over-production of stars at the bright end of the galaxy luminosity function by these models has led to the inclusion of an extra source of heating or “AGN feedback” (see e.g. Bower et al. 2006; Croton et al. 2006; Granato et al. 2004). However, it is still unclear what kind of AGN-driven feedback is the most important, e.g. two complementary mechanisms are described by Croton et al. (2006): “quasar-mode” feedback, proposed to be effective during episodes of efficient accretion of cold gas, and “radio-mode” feedback, caused by the relatively inefficient accretion of hot gas (in the absence of cold gas) when an AGN would have lower optical luminosity, but would also be generating radio jets. Thus, it is now clear that if we are to understand galaxy formation and evolution, we must also obtain a clear picture of AGN, their environments and the co-evolution of SMBHs and their host galaxies.

It is difficult to study the host galaxies of distant quasars, due to the nuclear quasar emission overwhelming the stellar emission at optical and near-infrared wavelengths. However, Hubble Space Telescope and high spatial-resolution ground-based imaging have enabled the investigation of the stellar populations of quasar hosts (e.g. Bahcall et al. 1994; McLure et al. 1999; Dunlop et al. 2003; Kotilainen et al. 2009), demonstrating that most luminous quasars reside in massive elliptical galaxies. A number of studies have attempted to determine the star-formation properties of quasar host galaxies, with some using optical colours (e.g. Sánchez et al. 2004) or spectroscopy (e.g. Nolan et al. 2001; Silverman et al. 2003; Trichas et al. 2010). Some information on the host galaxies of AGN has been gleaned from powerful radio galaxies, where the quasar nucleus is obscured due to the putative dusty torus (e.g. Antonucci 1993). Again these studies show that powerful AGN reside in massive galaxies (e.g. Jarvis et al. 2001; McLure et al. 2004; Seymour et al. 2007) and there appears to be a trend where the most powerful exhibit signs of recent star formation (e.g. Baldi & Capetti 2008; Herbert et al. 2010). However, optical colours and emission lines can be contaminated by the quasar, so that its contribution must be modelled to estimate the star-formation rate, and all of these studies suffer from the uncertainty arising from having very little information about how much star formation is obscured by dust, and
thus how much of the emission due to young stars is reprocessed to far-infrared and sub-millimetre wavelengths.

Spectroscopy with the *Spitzer Space Telescope* has enabled the determination of star-formation rates in individual AGN via mid-infrared emission lines, which are much less susceptible to dust, (e.g. Schweitzer et al. 2006; Netzer et al. 2007; Martínez-Sansigre et al. 2008; Lutz et al. 2008; Shi et al. 2004; Trichas et al. 2004; Lutz et al. 2010). However, the relatively low sensitivity of *Spitzer*s infrared spectrograph (IRS) has meant that only small, severely flux-limited samples of objects could be investigated, making the precise relationship between star-formation rate and quasar luminosity difficult to robustly determine.

(Sub)millimetre instruments such as SCUBA and MAMBO have provided important information on reprocessed emission from star formation in distant quasars and radio galaxies, e.g. photometric surveys revealing luminous ($\sim 10^{33} \text{l}_\odot$) far-infrared emission from a number of radio galaxies and quasar hosts, making it clear that major (SFR $\sim 1000 \text{M}_\odot \text{yr}^{-1}$) episodes of star formation can coexist with powerful AGN (e.g. Archibald et al. 2001; Priddey et al. 2003; Omont et al. 2003; Page et al. 2004; Martínez-Sansigre et al. 2009). Furthermore, sub-mm imaging has revealed overdensities of sources in the vicinity of individual high-z AGN, plausibly belonging to the same large-scale structures (Stevens et al. 2003; Priddey et al. 2002; Stevens et al. 2010). Conversely, cross-correlating SCUBA data with deep X-ray surveys, Alexander et al. (2003) also show a large fraction of sub-mm galaxies contain buried AGN, while Laird et al. (2010) find that the X-ray AGN fraction in sub-mm galaxies is 20–29%.

The main drawback of the sub-mm studies to date is that, due to the high background associated with ground-based observing and the small number of pixels in existing instruments, it has not been possible to obtain very large samples of objects. A recent study by Serjeant & Hatziminaoglou (2009, hereafter SH09) attempted to overcome this shortfall by combining data from a wealth of *Spitzer* and *ISO* observations at mid- and far-infrared wavelengths to investigate how obscured star formation in the hosts of quasars evolved over cosmic time and also how this may be related to the bolometric output of the AGN. They found evidence for an increase in far-infrared luminosity as a function of both cosmic epoch and quasar luminosity. However, they measured the correlation between far-infrared and quasar luminosity using wide redshift bins, which can potentially confuse the effect of redshift with that of quasar luminosity, since the latter evolves strongly with redshift in flux-density-limited samples. In addition, their work extrapolated the far-infrared flux from mid-infrared wavelengths using SED templates, and it would clearly be preferable to measure the far-infrared flux directly.

The recently-launched ESA *Herschel* space observatory (Pilbratt et al. 2010) offers an unprecedented combination of sensitive detectors, large (3.5-metre diameter) collecting area, and low background, which makes it possible, for the first time, to conduct deep surveys of large areas of the sky at far-infrared/sub-mm wavelengths.

In this paper we present a study of the far-infrared fluxes of known quasars in science demonstration observations of the *Herschel* Astrophysical Terahertz Large Area Survey (H-ATLAS; Eales et al. 2010). This is the first single sample of quasars measured in the sub-mm which is large and diverse enough to directly determine a relationship between far-infrared (cool dust) luminosity and optical (quasar) luminosity, without using infrared information at wavelengths short enough to be possibly contaminated by warm dust in the AGN torus. While Serjeant et al. (2010) determine a relationship between optical and far-infrared luminosity using this data in combination with infrared luminosities derived from shorter-wavelength *Spitzer* and *ISO* data, here we use only the H-ATLAS data.

Section 2 describes the H-ATLAS observations used. Section 3 explains how we construct our quasar sample, determine far-infrared fluxes for them, and find a maximum-likelihood fit to the relationship between far-infrared and optical luminosity, and redshift. The results of this fitting are discussed in Section 4, and conclusions are presented in Section 5.

## 2 OBSERVATIONS

When completed, H-ATLAS will provide photometry for 550 square degrees of the sky in five far-infrared/sub-mm bands. The 5-σ depths at wavelengths of 100, 160, 250, 350, and 500 μm are 132, 126, 32, 36 and 45 mJy respectively.

In this paper, we use only data from *Herschel*s SPIRE instrument (Griffin et al. 2010), at 250, 350, and 500 μm, since the PACS 100 and 160 μm data (Poglitsch et al. 2010) are not deep enough to provide useful constraints on the infrared luminosities of the quasars we are studying (Ibar et al. 2010). The longer SPIRE wavelengths are also less likely to be contaminated by direct emission from the warm dust of the AGN torus. This means that the emission can be more reliably attributed to cool dust heated by star-formation processes, and its spectral energy distribution (SED) approximated by a simple greybody.

This paper is also restricted in area to the H-ATLAS science demonstration (SD) field, which covers approximately 16 square degrees – around 3 per cent of the total survey area. This SD data has been publicly released, and both the maps and catalogues can be retrieved from the H-ATLAS webpage.

## 3 METHOD

Our quasar sample consists of all objects in the field which are identified as quasars in the seventh data release of the SDSS Quasar Catalog (Schneider et al. 2010), or the 2dF-SDSS LRG (luminous red galaxy) and Quasar spectroscopic catalogue (2SLAQ; Croom et al. 2009). After removing the known blazar HB0906+015, since its far-infrared emission will not be dominated by star formation (see e.g. González-Nuevo et al. 2010), we find 372

1 http://www.h-atlas.org
quasars in the whole SV field, with a reasonably conservative cut on the area (by eye) to avoid using the noisier data at the edges of the field.

Of the 372 quasars, only 29 have matches to objects in the 5-sigma H-ATLAS catalogue (Rigby et al. 2010; Smith et al. 2010). To improve our statistics, we measure fluxes directly from the SPIRE maps described by Pascale et al. (2010) at the position of each quasar. To do this, we take a point measurement of the flux in the nearest pixel to the SDSS position on each PSF-filtered and background-subtracted map.

To account for source confusion, the most likely flux from the object (which is used to produce Figure 2) is obtained by subtracting the mean pixel value in the map. However, when we perform a maximum-likelihood fit for the coefficients of Equation 2 as described below, there is an implicit assumption that the measurement noise is symmetric about zero. The confusion noise is very significantly skewed, and can cause false correlations to be found in random data, so for the fluxes used in these fits we symmetrise the noise distribution by subtracting the flux in a different randomly selected pixel from each flux measurement. This increases the overall noise level, so we average the likelihoods from a number of random background realisations in order to reduce the noise level in the final result.

We also construct a comparison sample of fluxes drawn from random positions in the maps, with their backgrounds subtracted in precisely the same way as the real fluxes, and are then used to produce maximum-likelihood estimates in the same way as the real data. Comparing these estimates with those from the real data allows us to increase our confidence that any trend detected in the data is due to the influence of the quasars, and not merely a result of the noise properties of the maps, which have a significant impact due to our inclusion of individual noisy objects instead of using binned data.

To compare the AGN luminosity of the quasars with that from star formation, we make the assumption that the optical light is dominated by the AGN, while the far-infrared/sub-mm is dominated by the cool dust heated by star formation. This latter assumption is consistent with the results of Hatziminaoglou et al. (2010), who fit models with AGN and star-formation contributions to the mid- and far-infrared emission from AGN hosts and find that star formation dominates the far-infrared fluxes. Their result depends on the assumption that the dusty torus heated directly by an AGN has a fairly simple, approximately toroidal geometry, as described by Fritz et al. (2006). The infrared SEDs of AGN can, however, be fully reproduced by more complex clumpy torus models (e.g. Nenkova et al. 2008), so, at least until these models are tested by high-spatial-resolution imaging of AGN tori, we cannot completely rule out the possibility of a contribution to far-infrared fluxes from dust heated directly by the AGN. However, the fact that Hatziminaoglou et al. (2010) find that the SPIRE colours of AGN are indistinguishable from those of normal galaxies suggests that the AGN does not make a strong direct contribution at these wavelengths.

All of the quasars have spectroscopic redshifts from their respective surveys, and optical photometry in the ugriz bands from the SDSS (York et al. 2000). We use these data to estimate the quasar luminosity, following Schneider et al. and others in using the absolute i-band magnitude, M_i, as a proxy for the luminosity of the central engine. This is calculated from the observed i magnitude (after the subtraction of Galactic extinction) and the redshift, by assuming a power-law SED (S_ν ∝ ν^α) with spectral index α = −0.5 and assuming a ΛCDM cosmology with Ω_m = 0.3, Ω_Λ = 0.7, and H_0 = 70 km s^{-1} Mpc^{-1}. Figure 1 shows the distribution of optical luminosities of our sample as a function of redshift, indicating the source of the spectroscopic redshift (SDSS or 2SLAQ, or both). From this we can see that the 2SLAQ measurements in this field allow the inclusion of quasars at least one magnitude fainter (at z < 3) than the SDSS spectroscopic survey probes. This greater range in optical luminosity is crucial for decoupling optical luminosity and redshift, and gives our study an advantage over similar work by Hatziminaoglou et al. (2010). We note, however, that the range of bolometric luminosities spanned by our quasar sample is 10^{45} < L_{bol} < 10^{48} erg s^{-1}, so all of the objects are firmly in the quasar (as opposed to Seyfert) regime. Thus, if there are two “modes” of accretion with different feedback properties, this study will only be relevant to understanding the “quasar mode”, making it usefully complementary to the work of Shao et al. (2010), which uses different Herschel data to probe the star-formation rates in lower-luminosity X-ray-selected AGN.

As mentioned earlier, the SPIRE photometry at 250, 350, and 500 μm is unlikely to be contaminated by the AGN directly for even the higher redshift quasars in the sample, and is deeper than the H-ATLAS PACS photom-

Figure 1. Absolute i-band magnitudes (calculated as described in the text) for our sample of spectroscopically confirmed quasars, plotted as a function of redshift. Objects with spectra from SDSS are shown as black diamonds; those with 2SLAQ spectra have red squares.
by integrating the best-fitting greybody with \( T = 30 \, \text{K} \) and \( \beta = 1.5 \) over the 8–1000 \( \mu \text{m} \) wavelength range) for our sample of spectroscopically confirmed quasars, plotted as a function of redshift. Objects with spectra from SDSS are shown as black diamonds; those with 2SLAQ spectra have red squares. Error bars are plotted where \( L_{IR} \) is greater than the uncertainty in the estimate, otherwise 1-\( \sigma \) upper limits are shown using arrows.

Figure 2. Estimated total infrared luminosities (calculated by integrating the best-fitting greybody with \( T = 30 \, \text{K} \) and \( \beta = 1.5 \) over the 8–1000 \( \mu \text{m} \) wavelength range) for our sample of spectroscopically confirmed quasars, plotted as a function of redshift. Objects with spectra from SDSS are shown as black diamonds; those with 2SLAQ spectra have red squares. Error bars are plotted where \( L_{IR} \) is greater than the uncertainty in the estimate, otherwise 1-\( \sigma \) upper limits are shown using arrows.

We estimate the uncertainty on \( L_{IR} \), so we use only the SPIRE fluxes to constrain the overall infrared luminosity. To make the results directly comparable with each other, we fit a simple isothermal grey dust model, with flux density

\[
S_\nu \propto \nu^{(3+\beta)}/(e^{h\nu/kT} - 1)
\]

where \( h \) and \( k \) are Planck’s and Boltzmann’s constants, \( \nu \) is rest-frame frequency, and where we assume fixed values for both temperature \( T = 30 \, \text{K} \) and \( \beta = 1.5 \). We then integrate the greybody curve between rest-frame wavelengths of 8 and 1000 \( \mu \text{m} \) to obtain an estimate of \( L_{IR} \).

We do not attempt to use a luminosity- or redshift-dependent dust temperature in this work. Increasing the dust temperature serves to shift the greybody peak to shorter wavelengths, which would result in a higher luminosity fit to our data (which are mostly longward of the peak). Thus, if the temperature of the dust in quasar hosts does increase with redshift and/or luminosity, then we will systematically underestimate \( L_{IR} \) for objects at high redshift and \( L_{IR} \). However, since the trend in mean temperature found by Amblard et al. (2010) is weak compared with the dispersion around the relation, we do not expect this to be a very significant effect.

Since we include far-infrared flux measurements at our quasar positions regardless of whether there is a significant detection in the far-infrared, fitting the normalisation of a fixed template is the only robust measurement possible. This procedure does not account for the possibility of dust temperature or emissivity evolution either with redshift or quasar luminosity, which may well be important effects to consider. However, we are able to use the well-detected objects in our sample to address the suitability of our choice of \( T \) and \( \beta \).

To check our choice of greybody temperature, we fit greybody SEDs to the SPIRE fluxes of individual quasars with good detections (signal-to-noise > 3) in all three SPIRE bands, with \( \beta \) fixed at 1.5 and \( T \) allowed to vary. The best-fitting temperatures for eight out of the nine well-detected quasars were very similar, with mean 32.4 K and a standard deviation of 2.8 K, while one outlier fits a temperature of 58.5\( \pm \)14.0 K. The eight quasars span a redshift range of 1.1 < \( z < 2.3 \) and a range in infrared luminosity (estimated by integrating the greybody between 8 and 1000 \( \mu \text{m} \)) of \( 6 \times 10^{47} < L_{IR}/L_{\odot} < 2 \times 10^{48} \). They are consistent with the temperatures of galaxies as a whole as determined by Amblard et al. (2010), who find a mean temperature of 28 \( \pm \) 8 K for galaxies detected at a significance of 3-\( \sigma \) in the H-ATLAS SDP field, and a mean temperature of 30 \( \pm \) 9 K for a compilation of galaxies in H-ATLAS and a number of other far-infrared surveys. This consistency is unsurprising, since Elbaz et al. (2010) also found (using deeper data from Herschel) that the dust temperatures of radio-quiet AGN hosts were not significantly different from the normal galaxy population. Thus we consider that a temperature \( T = 30 \, \text{K} \) is likely to be a reasonable choice for our quasar sample, but is not completely certain, and we return to this point later.

The outlying object has substantial uncertainties on its temperature, so we do not consider it an indication of a serious problem with our use of a single temperature for all objects. However, the fact that it is at a higher redshift (\( z = 3.4 \)) and slightly higher luminosity (\( L_{IR}/L_{\odot} = 3 \times 10^{48} \)) than the other objects means that it is broadly consistent with the evolution of dust temperature, to higher temperatures at higher redshifts and luminosities, seen in several previous studies (Chapman et al. 2005; Kovács et al. 2004; Michalowski et al. 2010; Amblard et al. 2010) and others.

We do not attempt to use a luminosity- or redshift-dependent dust temperature in this work. Increasing the dust temperature serves to shift the greybody peak to shorter wavelengths, which would result in a higher luminosity fit to our data (which are mostly longward of the peak). Thus, if the temperature of the dust in quasar hosts does increase with redshift and/or luminosity, then we will systematically underestimate \( L_{IR} \) for objects at high redshift and \( L_{IR} \). However, since the trend in mean temperature found by Amblard et al. (2010) is weak compared with the dispersion around the relation, we do not expect this to be a very significant effect.

To get an indication of the effect of temperature (and its uncertainty) on the parameters we derive, we perform fits to all objects using both \( T = 30 \, \text{K} \) and \( T = 45 \, \text{K} \), where we have selected 45 K because it has a similar peak wavelength to the average \( z > 4 \) quasar SED constructed by Priddey & McMahon (2001) (who find parameters of \( T = 41 \pm 5 \) with \( \beta = 1.95 \pm 0.3 \)). We also note that, while \( \beta \) may not in fact be 1.5, the luminosity estimates obtained using \( \beta = 2 \) (using the best-fitting temperature at this \( \beta \)) differ by less than 1 per cent, so we consider that the selection of a value of \( \beta \) is broadly degenerate with the choice of \( T \), at least for the purpose of estimating infrared luminosity from SPIRE fluxes.

Following SH04, we assume that the dependence of the infrared luminosity on the optical quasar luminosity and redshift can be expressed as power laws, i.e.

\[
L_{IR, \text{model}} = A L_{QSO}^\beta (1 + z)^\xi
\]
with intrinsic (Gaussian) dispersion in \( \theta \) and \( \zeta \) of \( \Delta \theta \) and \( \Delta \zeta \) respectively.

Unlike Serjeant & Hatziminaoglou, who stack far-infrared data to look for trends in bins of redshift and optical luminosity, we fit for the power law coefficients of equation (2), and the normalisation \( A \), by finding the peak in the likelihood distribution calculated over all data simultaneously. Our approach has the substantial disadvantage that variations in the power laws as a function of redshift or luminosity will be hidden from us. However, under the assumption that the values of \( \zeta \) and \( \theta \) are the same at all redshifts and luminosities, our method enables us to make the most sensitive possible determination of these parameters, since we do not lose information through binning. Additionally, we are immune to the potentially confusing effects of correlations within bins.

We incorporate the dispersions \( \Delta \theta \) and \( \Delta \zeta \) by combining them with the uncertainty \( \sigma_i \) on each data-point. The likelihood, \( \text{Prob}(\text{data}|\theta, \zeta, \Delta \theta, \Delta \zeta, A) \) at a given point in parameter space is then given by

\[
\text{Prob}(\text{data}|\theta, \zeta, \Delta \theta, \Delta \zeta, A) = \prod_i e^{-\frac{\chi_i^2}{2}} \frac{1}{2\pi (\sigma^2_\theta + \sigma^2_\theta, i + \sigma^2_\zeta, i)}
\]

where

\[
\chi_i^2 = \frac{(L\text{IR},\text{data},i - L\text{IR, model},i)^2}{\sigma^2_\theta + \sigma^2_\theta, i + \sigma^2_\zeta, i},
\]

\[
\sigma_\theta, i = L\text{IR, model},i \ln L\text{QSO}, i \Delta \theta, \text{ and }
\]

\[
\sigma_\zeta, i = L\text{IR, model},i \ln (1 + z) \Delta \zeta .
\]

The prior space on the parameters \( \theta, \zeta, \Delta \theta, \Delta \zeta, \) and \( A \), over which the likelihood was evaluated, was chosen variably to conform more than 99.999 per cent of the likelihood. The priors were linear in \( \theta, \zeta, \Delta \theta, \) and \( \Delta \zeta, \) and logarithmic in \( A, \) with the limits \( \Delta \theta \geq 0 \) and \( \Delta \zeta \geq 0. \)

4 RESULTS AND DISCUSSION

Figure 3 shows the results of the likelihood evaluations, marginalised over all parameters except the two shown in each panel. This is the result of averaging the likelihoods obtained in 17 random realisations of the background, as described in Section 2. Contours labelled 1-\( \sigma, 2-\sigma, 3-\sigma, \) and 4-\( \sigma \) enclose 68.3, 95.4, 99.73, and 99.994 per cent of the total likelihood. We show the results of fits assuming both \( T = 30 \) K and \( T = 45 \) K to give an indication of the way in which our uncertainty in \( T \) should contribute to the uncertainty in the derived parameters.

This figure shows that, assuming the dust in quasars has a temperature \( T = 30 \) K, there is good evidence for correlations of \( L\text{IR} \) with both \( L\text{QSO} \) and \( z \), since the maximum likelihood values of both \( \theta \) and \( \zeta \) are significantly non-zero, and that while higher temperatures can mimic the effect of the correlation with \( z \), the correlation with \( L\text{QSO} \) is robust to simple changes in the overall dust temperature. For the \( T = 30 \) K case, the best-fitting values of each parameter, with their 68.3 per cent confidence limits (marginalising over all other parameters in each case) are:

\[
\theta = 0.22 \pm 0.08 \quad \text{(7)}
\]

\[
\zeta = 1.6 \pm 0.4 \quad \text{(8)}
\]

\[
\Delta \theta = 0.08 \pm 0.08 \quad \text{(9)}
\]

\[
\Delta \zeta = 3.4 \pm 0.7 \quad \text{(10)}
\]

Although, as can be seen in the lower left panel of Figure 3, there is some degeneracy between \( \theta \) and \( \zeta \), our data exclude \( \theta = 0 \) at approximately the 2.5-\( \sigma \) level. The (flat) prior on \( \theta \) allowed it to take values between -1.5 and 1 (i.e. a much larger range than we show in Figure 3), so we are confident that this is not simply an artefact of reaching the edge of the prior range.

The other panels of Figure 3 show that there is not a strong degeneracy between any other parameters, so it is not possible to reproduce the effect of a correlation of \( L\text{IR} \) with either \( L\text{QSO} \) or \( z \) by simply increasing the intrinsic dispersion in the other parameter; indeed these panels suggest that approximately the same result would be obtained if dispersion in the power law indices was ignored altogether.

While the figure shows that the detected evolution of \( L\text{IR} \) with redshift is very strongly dependent on the assumed value of \( T \) (although a temperature as high as 45 K for the whole sample seems unlikely), the correlation with \( L\text{QSO} \) seems to be robust to this uncertainty. We find a slightly weaker index, \( \theta \), for the correlation with \( L\text{QSO} \) than found by SH09, whose average result \((\theta = 0.44 \pm 0.07)\) is shown in figure 3 as the point with errors bars. This may in part be due to the relatively large redshift bins (e.g. 0.5 \( \leq z \leq 1, 1 \leq z \leq 2 \)) which were used in SH09, since, because of the strong degeneracy between \( \theta \) and \( \zeta \), these could allow trends with redshift to affect objects within a single bin. However, the disagreement is not very significant.

In contrast, performing the same maximum-likelihood fit to data with far-infrared fluxes drawn from random positions in the maps (with the same position used at each wavelength, and the dust-fitting performed assuming a temperature of 30 K) we obtain the results shown in Figure 4. The important thing to note is that the best-fitting value of \( \theta \) is zero for the random flux case, indicating that the result we have obtained is not simply due to the noise properties of the far-infrared maps, but rather is due to far-infrared flux genuinely associated with the quasars. (Note that the non-zero value for \( \zeta \) derived from the fit to random data is not unexpected, since the real quasar luminosities and redshifts were still used in this case, and these are themselves correlated.)

Using deeper far-infrared data from Herschel, but with a narrower range in optical luminosity (because only SDSS, and not 2SLAQ, quasars are available in the field), Serjeant et al. (2010) find a very similar result, \( \theta = 0.35 \), for objects at redshifts \( z > 2 \). Serjeant et al. (2010) use a compilation of H-ATLAS and other data to find a weakly declining correlation between \( L\text{IR} \) and quasar luminosity, parametrised as \( \theta = (0.5875 \pm 0.045) - (0.09275 \pm 0.012) z \) for redshifts \( 0 < z < 4 \), which is a slightly higher value of \( \theta \) than we find, but broadly consistent within the errors and at the redshifts we are sensitive to. Our value of \( \theta \) is also consis-

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Figure 3. Likelihood contours for pairs of the parameters $\theta$ and $\zeta$ (from equation \ref{eq:2}) and their intrinsic dispersions $\Delta \theta$ and $\Delta \zeta$, marginalised over the parameters not displayed in each panel. Contours labelled 1-$\sigma$, 2-$\sigma$, 3-$\sigma$, and 4-$\sigma$ enclose 68.3, 95.4, 99.73, and 99.994 per cent of the total likelihood. The point with error bars in the lower left-hand panel shows the average values of $\theta$ and $\zeta$, and their 1-$\sigma$ uncertainties, determined by SH09 via stacking in the mid-infrared. Solid contours show a fit to objects assuming $T = 30$ K; dotted contours show a fit to the objects assuming $T = 45$ K.

Tent with the work by Silverman et al. (2009), who find $\theta = 0.28 \pm 0.22$ for the slope of the power law relating [OII] luminosity (tracing star formation after correcting for AGN contamination) to X-ray luminosity (tracing the AGN). Silverman et al. study a sample of AGN over a narrower range of redshifts ($0.5 < z < 1$) and with somewhat lower luminosities than most of our sample ($10^{44.3}$ erg s$^{-1} \lesssim L_{\text{bol}} \lesssim 10^{45.8}$ erg s$^{-1}$, assuming a bolometric correction $L_{\text{bol}}/L_{2-10\text{keV}} \sim 30$); the consistency of their result with ours suggests that the relationship between $L_{\text{IR}}$ and $L_{\text{QSO}}$ may not vary with $L_{\text{QSO}}$ as suggested by Lutz et al. (2010) and Shao et al. (2010). Their steeper power law, with a slope of 0.8, for luminous quasars is not directly comparable with our result as it does not separate out the independent effect of redshift on $L_{\text{IR}}$. However, we would also note that the quasar part of this relationship is based on samples presented by Netzer (2009 and earlier works), which ignore AGN without detections in the far-infrared or mm, so that the observed correlation may be partly due to far-infrared flux limits.

To investigate whether our fit could be affected by a redshift-dependence in $\theta$, we perform the maximum likelihood fitting again, using objects in restricted redshift ranges where we have enough objects to make a robust measurement, and marginalise over all parameters except $\theta$. The results of this are shown in Figure 5; the derived
values of \(\theta\) are not significantly different from each other for redshifts \(0.5 < z < 2.5\), although the small number of objects in the redshift bins means we cannot make a strong statement for or against evolution of \(\theta\) as a function of redshift. However, the \(0.5 < z < 1\) range shows a detection of non-zero \(\theta\) at better than the 2-\(\sigma\) level, which is complementary to the result of Hatziminaoglou et al. (2010). While we do not have enough data to detect any significant decline in \(\theta\) as a function of redshift, the data-points in Figure 5 are consistent with the relation found by Serjeant et al. (2010), and with our smaller H-ATLAS-only dataset we cannot rule out the existence of such a weak redshift dependence.

Since the far-infrared luminosity \(L_{\text{IR}}\) has been shown to be due to star formation (Hatziminaoglou et al. 2010), the physical interpretation of these results is that star-formation rate is higher both as redshift increases and as quasar luminosity increases. As was also suggested by SH09, and as has been incorporated into the “quasar-mode” feedback recipes of semi-analytic models (e.g. Croton et al. 2006), one of the simplest ways to interpret this is if both black hole accretion and star formation depend on a common supply of cold gas, as might be supplied during a galaxy–galaxy merger or cold-accretion event (Dekel et al. 2009; Dijkstra & Loeb 2009; Smith & Jarvis 2007; Smith et al. 2008). The implication of the non-unity-slope is that the dependence of either star-formation or accretion luminosity (or both) on the gas supply cannot simply be linear with gas mass, as one might naively suppose, but is instead more complex.

For both star formation and black-hole accretion, such gas must be sufficiently cool that pressure support...
does not prevent it from being accreted by the AGN or from collapsing under its own gravity to form stars, so an alternative way to make this connection is via a mechanism (such as AGN or supernova feedback) which is capable of heating gas at all scales within a galaxy.

5 CONCLUSIONS

Using far-infrared measurements (at 250, 350 and 500 \(\mu\)m) of SDSS and 2SLAQ quasars in the H-ATLAS science demonstration field, we find evidence for correlations between the infrared luminosities of quasar hosts with both redshift and quasar luminosity. While the best-fit power-law index for the correlation with redshift can vary significantly depending on the assumed temperature of the dust in the quasar hosts, the correlation with quasar luminosity is robust to this effect, and is non-zero at the 2.5-\(\sigma\) level when we combine all objects in a maximum-likelihood fit.

We find a power-law index \(\theta = 0.22 \pm 0.08\) for the relationship between \(L_{\text{IR}}\) and \(L_{\text{qso}}\). Our data show that \(\theta\) is non-zero at the 2-\(\sigma\) level in the \(0.5 < z < 1\) redshift range alone, but we have too few objects to find evidence for evolution in \(\theta\) over the range \(0.5 < z < 2.5\).

We find evidence for a large intrinsic dispersion in the redshift dependence, but no evidence for intrinsic dispersion in the correlation between \(L_{\text{QSO}}\) and \(L_{\text{IR}}\), suggesting that the latter is due to a direct physical connection between star formation and black hole accretion. One possible interpretation of this is that both the quasar activity and star formation are dependent on the same reservoir of cold gas, and are thus both affected by influx of gas during mergers or cold accretion, or heating of gas via feedback processes.

With the full 550 square degrees of the H-ATLAS survey, we will be able to substantially increase the sample of quasars and improve these constraints. The larger area will allow us to obtain representative samples of the objects with the lowest density on the sky, i.e. quasars with high luminosities at low redshift, which will increase the sensitivity of our test at low redshifts and permit us to remove the assumption of power-law correlations and test more general models. More importantly, a larger sample of objects at all redshifts will make it possible to measure the temperatures of dust in quasar hosts as a function of \(L_{\text{qso}}\) and redshift by stacking the far-infrared fluxes; this in turn will remove the largest source of uncertainty in the relationship between \(L_{\text{IR}}\) and \(z\).

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